UNIVERSITA' CATTOLICA DEL SACRO CUORE - Milano -

## QUADERNI DELL'ISTITUTO DI ECONOMIA DELL'IMPRESA E DEL LAVORO

| The effects of high school choices on academic performance <br> and early labour market outcomes |
| :---: |
| Lorenzo Cappellari |
| n. 35 - aprile 2003 |

## QUADERNI DELL'ISTITUTO DI

ECONOMIA DELL'IMPRESA E DEL LAVORO
Istituto di Economia dell'Impresa e del Lavoro
Facoltà di Economia
Università Cattolica del Sacro Cuore
Largo Gemelli, 1-20123 Milano

# The effects of high school choices on academic performance and early labour market outcomes ${ }^{\text {b/ }}$ 

Lorenzo Cappellari<br>Dipartimento di Scienze Economiche e Metodi Quantitativi<br>Università del Piemonte Orientale<br>Via Perrone 18, 28100, Novara, Italy<br>Email: lorenzo.cappellari@eco.unipmn.it

April 2003


#### Abstract

Using microdata on the 1995 cohort of Italian high school graduates, this paper studies the relationship between the type of high school attended (general versus technical/professional; private versus public) and indicators of subsequent performance. Simultaneity issues that potentially bias this type of exercise are tackled by instrumental variables in a multivariate probabilistic framework. Results indicate that choices greatly depend upon the family of origin and prior school performance. General high schools are found to increase the probability of transition to university and to improve performance once at university. On the other hand, private high schools appear to be associated with lower academic performance. Technical schools improve the quality of the school-to-work transition, both in terms of employment probabilities and earnings.


Keywords: high school choices; academic performance; early labour market outcomes JEL codes: I21, J24, C35

[^0]
## 1. Introduction

Italian educational policies are the subject of an intense debate. In March 2003 the Italian parliament approved a law of reform of the high school system which introduces a new vocational track that includes class work in the first two years and a mixture of class work and on-the-job training in the last couple of years, plus an additional fifth year for those wishing to enrol into university. The vocational track will parallel the ones of general and technical/professional high schools, which are also reformed to allow the possibility of on-the-job training spells. The new law warrants the possibility of transition between tracks. ${ }^{1}$ Proponents of the new system emphasise the role that labour demand will play, via training schemes, in shaping the type of skills that vocational track students would acquire, therefore facilitating their school-to-work transitions; opponents stress the risks of social segmentation that such system could ingenerate through tracks separation, doubt the feasibility of transitions between tracks, and question the usefulness of learning job-specific skills during school years in a world of fast skill obsolescence.

A second issue debated is school provision. Since 2000, school vouchers started being introduced by some regional governments, with the aim of helping households in choosing their preferred type of school, whether private or public. Given the difference in costs entailed by the two types of provision, vouchers lower the relative price of private schools in terms of public ones, thereby shifting some demand for schooling from the latter to the former. In conjunction, decentralization of schools organisation is also an item on the government's agenda, and vouchers can be expected to spread to other regions as regional governments gain power on school issues. Both freedom of choice and increases in private and public school quality (the latter brought about by greater competition) can be found among the principles inspiring vouchers programmes. Opponents to these policies question the ability of private schools in providing quality education and argue for employing public funds for improving the effectiveness of public schools.

The policy debate prompts several scientific questions. To what extent the choice of a high school track affects later individual outcomes in terms of access to higher education, academic performance, and the school-to-work transition? How does the performance of private school students compare with that of public schools ones? Do differences in performance across high school types reflect a causal effect or are they due to endogenous
sorting of students across school types? This paper aims at providing answers to the above questions. Using survey data on a cohort of high school graduates, models of the impact of high school types (general versus technical/professional; private versus public) on the transition to university, academic performance, labour market participation, employment probabilities and wages are estimated. In order to unravel the causal link between school choices and subsequent outcomes, great attention is devoted to endogeneity issues, which are tackled by employing instrumental variables within a multivariate probabilistic framework.

The literature on the evaluation of school type effects on measures of performance has grown rapidly over the past decade. For example, the impact of catholic schools on academic performance has received considerable attention in the United States. Evans and Schwab (1995) highlight the endogeneity issues entailed by this type of exercise and use instrumental variables to identify the effect of catholic school attendance on measures of academic success, concluding that catholic schools raise subsequent educational outcomes. Neal (1997) uses area-level measures of catholic schools availability as instruments for school choices, showing that the benefit of catholic schools are confined to urban minorities, possibly as a consequence of the low quality of available public schools. An instrumental variable procedure is employed by Figlio and Stone (1999) to assess the effect of religious and non-religious private schools on educational outcomes, finding that, in general, only the former increase individual outputs relative to public schools. The methodological approach developed by Altonji et al. (2000) is instead based on the use observable information as a way of reducing endogeneity bias; they find that catholic schools are effective in favouring high school completion, while the effect on transition to college. Recently, researchers' interest on the effect of school choices have also started spreading onto other spheres of human life: an example is Figlio and Ludwig (2000), who look at the effect of catholic school attendance on youths' crime, drug abuse and sexual activity, finding that catholic schools are effective in reducing all three.

While catholic/private versus public schools effects have been extensively researched, the general versus technical or vocational high school divide is less explored, perhaps not surprisingly, given that US high schools are characterised by greater uniformity under this respect compared other countries, Italy included. For France, Margolis and Simonnet

[^1](2002) show that technical high school graduates outperform general high schools ones in the school-to-work transition, thanks to the more effective labour market networks they can access. Positive earnings effects of vocational high schools compared to general high schools are reported by Moenjak and Worsick (2003) on a sample drawn from the labour force survey of Thailand, after correcting for endogenous selection issues.

Given their current political relevance, it is not surprising that the interest surrounding economics of education issues in Italy is increasing at a fast rate. The main issue emerging from the Italian literature is a strong intergenerational persistence in educational achievement. Checchi et al. (1999) contrast intergenerational mobility in incomes and schooling attainment between Italy and the US, finding larger persistence in the former case, and argue that the Italian schooling system might have failed to provide poor families with the incentives to invest in the human capital of their off-springs. Besides family of origin, also high school types appear to be strongly associated to educational outcomes. Bertola and Checchi (2002) study a sample of university students from the University of Milan and find that those coming from general and public high schools score better than otherwise comparable students on a range of performance indicators. They also consider the differences in academic performance between public and private school students, finding that public schools are associated to the better performances, followed by religious private schools and lay private schools. The issue of public/private school choices is analysed in Checchi and Jappelli (2002) using subjective assessment of public school quality: they do not find evidence of any quality differential in favour of private school as determinant of the school choice. A theoretical perspective on the optimal school design is provided by Brunello and Giannini (2000), showing that considerations on the desirability of educational stratification can not be unambiguously drawn from an efficiency point of view.

This paper's empirical strategy extends the bivariate probit-type of models employed by Evans and Schwab (1995) and Neal (1997) among others. From a substantive point of view, these extensions allow to control for two endogenous treatments simultaneously (general versus technical/professional high school; private versus public high school) and for the presence of endogenous censoring on some of the outcomes of interest. Results show how these extensions are relevant, results changing substantially when moving from completely exogenous models to the models introduced in this paper. Estimates show that

[^2]choices greatly depend upon the family of origin and prior school performance. General high schools are found to increase the probability of transition to university and to improve performance once at university. On the other hand, private high schools appear to be detrimental for academic performance. Technical schools improve the quality of the school-to-work transition, both in terms of employment probabilities and earnings.

## 2. Data and descriptive patterns of high school choices and subsequent performances

The data used in this paper originate from the "1998 Survey on the school and work experiences of 1995 high school graduates", a cross-sectional sample of 18,843 high school leavers interviewed by the National Statistical Office (ISTAT) three years after graduation. The sample represents approximately 4 percent of the population of Italian high school graduates of 1995 and contains a wide range of information on the high school curriculum and on post high school experiences, either in the tertiary education system - including university - and the labour market. In addition, information on personal characteristics and family background is available.

The Italian high school system at the time of interview may be broadly described by three types of schools: general ('licei'), technical/professional ('istituti tecnici e professionali') and teaching schools ('istituti magistrali'). The first two types are based on a five-years curriculum, at the end of which students can freely chose to enrol in the university. ${ }^{2}$ The third type is specifically aimed at training elementary school teachers and is structured into a four years curriculum, plus an additional year required to students who wish to go to university. In the present paper this latter group (which represented approximately 12 percent of the sample originally available) has been excluded from the estimation sample, due to the difference in the structure of the curriculum compared to the other two types. Thence, the choice of the high school track analysed here focuses on the general versus technical/professional divide.

As for the other dimension of the Italian debate on educational policies, i.e. the public versus private provision of schools, the Survey enables identification of the type of school attended by asking respondent if the school they enrolled in after compulsory education was private or public, and if they made any transition from private to public school (and

[^3]vice versa) during high school. Differently from the work of Bertola and Checchi (2002), the data do not contain information on the confessional nature of private schools (see

The distribution of the type of school attended (general versus technical and public versus private) in the estimation sample is reported in Table 1. The estimation sample also excludes those who are employed and started their job while at high school (3.6 percent of the original sample), since their post-graduation choices might not be comparable with those of the rest of the sample. In addition, the estimation sample for the analysis of academic performance excludes students who enrolled into short university diplomas (2.2 percent of the original sample) because it was not possible to compute reliable indicators of academic performance due to small cells (see below for the definition of performance used), and those who did not report the information necessary for computing academic performance ( 0.05 percent). Graduates from general high schools represent 31 percent of the sample, whereas 12 percent of it graduate from a private high school. The two dimension of high school choices are strongly correlated. The probability of having a private school diploma is 9 percent among graduates from technical schools and 19 percent among students from the licei. The other conditional frequencies presented in the table show that the incidence of students from general high schools is some 17 percent larger among private school graduates than in the overall sample. One explanation of the positive association between the two variables can be found in the larger supply of private education in the high schools system: official statistics for 1989 indicate that while 14 percent of technical schools is private, the proportion rise to 30 percent when general high schools are considered. A second reason has to do with family resources: both private and general schools entail larger costs than public or technical ones, either in term of fees or expected opportunity costs.

Household incomes are not available in the data; however, the survey contains rather detailed information on parental education and occupation which partly compensates for the unavailability of incomes. The association between school types and parents' education is described in Table 2. The more educated any of the parents, the more likely it is that the children have general type of degree, with the probability that roughly quadruples when going from parents with none or low educational attainment to parents with university degrees. As long as education is positively associated with incomes, this finding might reflect larger financial endowments of high-education families, which can thus afford to place their children in a track that is more likely to continue with university compared to
the technical track. A second explanation has to do with preferences, as long as more educated parents attach higher value to education then less educated ones, and thus prefer general tracks which are more likely to continue with higher education. Finally, parents education might influence children studying capabilities, thereby lowering the costs of their investments in education. Parents' education is also positively associated with the probability of graduating from private schools, possibly reflecting financial endowments.

The association between parents' occupation and school choices is described by Table 3. General school graduates typically tend to come from families where parents are in highlevel non manual, managerial or professional occupations. Moreover, also teachers' children tend to graduate in general schools. On the other hand, private school leavers are likely to come from families where parents are self-employed.

The link between previous academic performance (measured by the mark reported in the final exam of the junior high school) and school choices is illustrated in Table 4. The better the performance, the larger the probability of graduating from general schools, rising by 7 times when moving from the lowest to the highest mark. Differences in ability across marks are an explanation for this evidence, the other possibility being better performances from children from high education/high income families compared to the rest of the sample. The opposite link, i.e. decreasing with the marks, can be observed for the probability of graduating in private schools, although the variation is much less astonishing than in the case of the probability of graduating in general schools. The negative association might signal that private education is used as remedial education, and is therefore more frequent among students at higher risk of experiencing problems in the course of their academic career (see Bertola and Checchi, 2002, on this point).

After having described the associations between family backgrounds and high school choices, in Table 5 the links between choices and subsequent academic performance are considered. Two are the dimensions of academic performance taken into account. The first is the probability of attending university at the time of interview. The second is the speed in passing university exams. In particular, for each of eight broad disciplinary areas in which university courses can be grouped (sciences, medicine, engineering, architecture, economics, business administration and statistics, political sciences, law, and the humanities), the median number of exams passed per year has been computed; the speed indicator is a dummy for being above the median in the distribution of the numbers of exams passed per year. Clearly, this is an imperfect performance indicator, since it does not
take the marks obtained into account. However, no information on marks is reported in the survey. Moreover, one of the main concerns inspiring the recent reforms of university degrees has been the long duration of university studies in the old system, so that studying the impact of high school choices on the speed in taking exams will yield insights on an issue at the core of Italian educational policies.

Table 5 shows that 40 percent of high school graduates attend university three years after graduation. The proportion halves among technical school graduates and doubles among general school graduates, revealing a strong link between type of degree held and transition to university. Transitions to university are more frequent among private school graduates than they are among public schools ones, although the magnitude of the link is less evident than in the previous case. Due to discontinuities in the distribution of exam passed per year, the sample proportion of cases above the mean (the 'fast takers') is 45 percent. Graduating from a general high school changes this proportion only slightly, by 4 percent; on the other hand, the proportion of fast exam takers among technical school graduates is 38 percent. Evidence on the private/public school divide shows little variation in speed.

The second dimension of students performance considered here are early labour market outcomes; in particular, the probability of engaging in job search activities, the probability of being employed and the probability of being low paid are taken into account. The ISTAT questionnaire contains a wide array of information on both job search and the characteristics of the current job. As for academic performance, also in this case graduates from teaching schools and those who started their current jobs while at high school have been excluded from the sample. In addition, after testing for endogenous selection and failing to reject the null hypothesis of exogeneity, those who work as self-employed (4.6 percent of the original sample) have been excluded from the estimation sample. Also, excluded are employees with missing wages (an additional 2.6 percent of the original sample). Therefore, the estimation sample for the analysis of economic performance differs from the one used for analysing academic performance, and consists of 14,420 observations. The proportions of graduates from general and private schools, as well as the variation of these proportions by family backgrounds in this sample are very similar to the ones shown in Tables 1-5.

Table 6 illustrates the variation of early labour market outcomes in the sample. An individual is defined as having done some search if she has an occupation of any kind or if
she reports herself being on job search. The employment status considered refers to continual employment; unemployed individuals having done some work in the week prior to interview or those on occasional or seasonal jobs are not counted as employed. Partly, such a choice is driven by the data, since no information on job attributes is collected for seasonal or occasional employment; in addition, it allows to focus on stable employment, which is probably more relevant from a policy perspective. Low pay is defined as the bottom quartile of the sample distribution of net hourly pay, i.e. the low pay is defined relatively to the group of high school graduates. After graduation, job search occurs for 68 percent of the sample. The proportion rises to 81 percent among technical school graduates and falls to 40 percent for general school ones, clearly indicating a difference in the propensity for labour market activities between the two educational tracks. A slight difference in job search propensities can also be observed between public and private schools graduates. Of those who search, 42 percent are observed in employment. This rather low figure depends upon the definition of employment adopted: if one included also seasonal employment, occasional employment and the unemployed reporting hours of works in the week prior to interview, the employment rate would be 68 percent. The employment rate varies depending upon the type of high school attended, passing from the 49 percent of technical schools graduates to the 16 percent of licei students; some variability is evident also between private and public schools. Finally, some associations between school choices and the probability of earning low wages can be detected, although they are less evident compared to the case of employment.

## 3. High school choices and academic performance

The previous section has shown that the link between the type of high school diploma and the probability of going to university is well evident in the data, and similar conclusions apply for the speed in passing exams once at university. In order to disentangle the effect of school types on outcomes, endogeneity issues that are likely to influence raw correlations have to be tackled. One first source of endogeneity has to do with selection into school types, as long as unobserved ability and unobserved parental background might influence both choices and outcomes. Second, the analysis of exams speed can only be performed conditional on university attendance; as long as the sample of university students is not representative of the population of high school graduates, estimates can be biased. What is needed, therefore, is an econometric model that jointly handles endogenous school choices
and endogenous sample selection. Given that the parameters of interest are unconditional on labour market decisions (i.e. estimating the effect of job search decisions on university attendance - or viceversa- is not the aim of this paper) the academic performance model is estimated without taking those decisions into account. Robustness checks performed by estimating high school choices, university attendance and job search probabilities simultaneously (without conditioning university and search equations upon each other), showed that while the reduced form correlation between job search and university attendance is negative and very precisely estimated, results on the effects of high school choices in the two equations - or indeed the coefficients associated to other covariates in those equations-are hardly affected by the simultaneous modelling of the two processes. Results from these checks are available upon request.

### 3.1 The model of academic performance

In this paper the endogeneity issues discussed above are assessed by means of multivariate probit models. Let observations in the estimation sample be indexed by $i, i=1 \ldots N$, and let the latent propensities to graduate from general schools $\left(g^{*}\right)$ and private schools $\left(p^{*}{ }_{i}\right)$ be linear functions of individual characteristics:

$$
\begin{align*}
& g^{*}=\beta_{g}{ }^{\prime} x_{g i}+\varepsilon_{g i}  \tag{1}\\
& p^{*}=\beta_{p}{ }^{\prime} x_{p i}+\varepsilon_{p i} \tag{2}
\end{align*}
$$

where the $x$ s are vectors of observable attributes such as gender, parental background and indicators of academic performance at the junior high stage, the $\beta \mathrm{s}$ associated parameter vectors to be estimated and the $\varepsilon s$ error terms assumed to be distributed as standard normal. When the latent propensities cross some unobserved thresholds -which can be fixed at zero without loss of generality- individuals are observed to graduate from general or private schools; let $g_{i}=\mathrm{I}\left(g_{i}^{*}>0\right)$ and $p_{i}=\mathrm{I}\left(p_{i}{ }_{i}>0\right)$ be dummy variables indicating the two events, the indicator function I( ) taking value one whenever its argument is true and zero otherwise.

Next, assume that the latent propensity to be observed at university at the date of interview, $u^{*}$, depends linearly on a set of personal characteristics $z_{u}$, and high school choices $g_{i}$ and $p_{i}$ :

$$
\begin{equation*}
u_{i}^{*}=\beta_{u} z_{u i}+\delta_{g} g_{i}+\delta_{p} p_{i}+\varepsilon_{u i} \tag{3}
\end{equation*}
$$

where the error term $\varepsilon_{u}$ is assumed to be distributed as standard normal; let $u_{i}=\mathrm{I}\left(l_{i}>0\right)$ be a dummy variable signalling university attendance at the date of interview.

Finally, conditionally on university attendance, academic performance can be observed. As explained in Section 2, speed in passing exams is the performance indicator available in the data. Let $n_{i}$, the average number of exams passed per year by individual $i$, depend on personal attributes $m_{s}$ and high school choices according to the following relationship:

$$
\begin{equation*}
s^{*}{ }_{i}=h\left(n_{i}\right)=\beta_{s} m_{s i}+\gamma_{g} g_{i}+\gamma_{p} p_{i}+\varepsilon_{s i} \quad \text { if } u_{i}=1 \tag{4}
\end{equation*}
$$

where $h()$ is a monotonic unspecified transformation such as the error tem of (4) is distributed as standard normal, while the relationship is unobservable when individuals are not at university, i.e. when $u_{i}=0$. An individual is classified to be a 'fast taker' whenever she lies in the upper half of the subject-specific distribution of $n_{i}$, let $s_{i}=\mathrm{I}\left(n_{i}>\mu\right)$ signal that event, where $\mu$ is the median of the subject-specific distribution of exam passed per year. Note that speed is conditional on both high school choices and university attendance.

The vector of error terms $\varepsilon_{i}^{A}=\left(\varepsilon_{g i}, \varepsilon_{p i}, \varepsilon_{u i}, \varepsilon_{s i}\right)$ is assumed to follow the four-variate normal distribution:

$$
\begin{equation*}
\varepsilon_{i}^{A} \sim \aleph_{4}(0, \Omega) \tag{5}
\end{equation*}
$$

where the correlation matrix $\Omega$ has unit diagonal elements and extra-diagonal elements equal to $\rho_{j k}, j, k=g, p, u, s$. Overall, the model specified is a four-variate probit with censoring of one equation. ${ }^{3}$

Estimation of cross-equation correlation coefficients allows to control for unobserved heterogeneity, therefore eliminating the issues of endogeneity discussed above. The potential endogeneity of high school choices and selection into university can be tested by testing the significance of the elements of $\Omega$.

[^4]Model identification requires valid "instruments", i.e. variables that can be excluded from the outcome equations ( $u^{*}$ or $s^{*}$ ) but significantly affect high school choices or selection into university. Candidates as instruments for high school choices are measures of family characteristics relevant at the time the choice was made, but whose effect on students behaviour can be expected to be less relevant in the university years, after their impact on high school choices has been taken into account. The instrument used for general high school is a dummy for having a grandparent with a university degree. The private school equation includes the same variable and a dummy for the mother being retired when the individual was $14 .{ }^{4}$ Instrumenting selection into university requires variables that affect the choice of going to university but have no residual impact on exams speed. In particular, dummies for the number of siblings in the household of origin will be used here.

### 3.2 Results

Table 7 presents in column (1) results from the estimation of the four-variate probit with censoring; in particular, it provides "marginal effects", evaluated at the means of explanatory variables. For comparative purposes, column (2) of the table also reports estimation of academic performance equations from univariate probit models, i.e. ignoring the endogeneity issues.

The bottom part of the table shows results from tests of significance of the instruments, performed using functional form as identifying restriction. The tests indicate that the data support the choice of instruments made: high school choices are significantly affected by grandparents' education and by the mother retirement status when the individual was 14 , but have no residual impact on the transition to university; similarly, the presence and number of siblings affects the probability of attending university at the date of interview, but have no impact on exam taking speed.

Three of the cross-equation correlation coefficients are statistically significant at conventional confidence levels, whereas in the other three are not statistically different from zero. The correlation between general high schools and private high schools ( $\rho_{g p}$ ) appears to be positive and very precisely estimated, confirming the evidence already emerged in Section 2 about the strong link between the two variables. Differently from there, however, in this section controls are made for family resources, so it is likely that the

[^5]positive coefficient reflects a larger supply of private education among general schools than among technical ones. The correlation between private high schools and exams speed ( $\rho_{p s}$ ) is also positive, indicating that the group of private school graduate that go to university is faster in taking exams compared to the average of university students. Finally, the positive estimate of the correlation between university attendance and exams speed ( $\rho_{u s}$ ) indicates that there is positive selection into university, i.e. those who are not observed at university would have been slower in taking exams compared to those currently attending; if exams speed can be deemed as an ability indicator, this result suggests that those observed at university are intrinsically more able than those not attending, net of family backgrounds. The table also reports tests for the overall ignorability of high school equations, i.e. tests for the hypotheses that all correlation coefficients referring to a certain high school equation are simultaneously equal to zero, and shows that both processes are non-ignorable for the estimation of academic performance. A test of overall significance of the correlation structure of unobservables also points to the necessity of tackling the various endogeneity issues.

The first two sub-columns of column (1) (P(general) and $\mathrm{P}($ private $)$ ) provide estimates of the determinants of high school choices. The presence of a favourable family background affects both choices significantly. In particular, while general high school graduates tend to come from families where one or both parents, are in managerial or professional occupations, private schools are chosen more frequently when the parents are self-employed or entrepreneurs. Also, teachers' children are likely to graduate from general and public high schools. Parental education positively and strongly influences both probabilities. The indicator of academic performance at junior high school, i.e. before entering high school, is strongly and positively associated to the probability of graduating from a general high school. Much smaller effects, and in the opposite direction, can be observed for private high schools. Overall, results appear to confirm previous findings from Bertola and Checchi (2002) on the fact that private schools attract students from richer families and with relatively lower ability compared to public schools ones, whereas general high schools attract higher performance individuals from richer families.

The estimated effects of graduating from a general high schools on the probability of attending university at the interview date is $43 \%$ for an individual with other observable attributes equal to the sample mean. This effect compares with an estimate of $47 \%$
retrievable from the model without endogeneity controls (see column (2), P (university)). The small change in the estimated effect between the two cases reflects the fact that the correlation of unobservables between the two equations $\left(\rho_{g u}\right)$ is not statistically different from zero. Note, however, that a dramatic drop characterises the asymptotic t-ratio when moving from the exogenous to the endogenous model, reflecting the fact that in the latter case the information is used over a wider parameter space. Similar evidence characterises the marginal effect for private school. In this case, the drop in significance is crucial, since it leads to conclude that there is no effect of private high school attendance on transitions to university. Thus, the significance of the negative effect emerging from the exogenous model is a consequence of omitting to control for the role of observable attributes - namely parental background and junior high school marks - in the determination of private school choices.

Estimates associated to parental education indicate a strong intergenerational persistence in higher education: the effect on university attendance roughly doubles when parental educational attainment passes from high school degrees to university degrees. Parents' occupation also affects university attendance significantly. A similar remark applies to junior high school marks. The regression also controls for high school marks and, not surprisingly, it appears to be strongly associated to the probability of being at university.

The importance of controlling for endogeneity is confirmed by the estimates of the exam speed equation. First of all, remember that the estimated correlation of unobservables between the university attendance and exam speed equations, i.e. the sub-sample on which speed is observed has an intrinsic speed larger that the one of the overall sample. This fact is reflected in the predicted probability of belonging to the upper half of the speed distribution, which is lower in the endogenous than in the exogenous model since in the former case it refers to the entire sample, which has a lower intrinsic speed than the population of university students. Another symptom of the allowance for endogenous selection is the rise in size of the effect of general high schools (remarks similar to those put forward when commenting P (university) apply to the drop in significance): general high schools have a positive impact on speed, and when such impact is computed in the entire sample, rather than from a high-speed sub-sample, the magnitude of the effect rises by more than $50 \%$. If factors that positively affect speed are magnified by the endogenous selection control, the opposite should hold for speed reducing variables: see, for example,
the drop in the marginal effect associated to residence in southern regions. In the case of the negative effects of private schools on speed, however, such an overestimation is counteracted by unobserved heterogeneity. The estimated positive correlation between unobservables of the private school and speed equation $\left(\rho_{p s}\right)$ indicates that individuals from private schools going to university have a speed propensity larger than the average. In the exogenous model the marginal effect of private schools on speed is a convolution of a negative causal effect and a positive selection term, which is removed from the marginal effect in the exogenous model. Estimates indicate that endogeneity is relevant, the negative effect of private schools on speed more than doubling when moving from the exogenous to the endogenous model.

Estimates of marginal effects on the other covariates in the speed equation point towards the importance of both parental backgrounds and individual ability (measured by marks reported at junior high and high schools) as explanatory factors of this dimension of academic performance.

## 4. High school choices and early labour market outcomes

The raw correlations presented in Section 2 point towards the existence of a strong link between the type of high school attended and behaviours at labour market entry; in this section the extent to which those associations can be imputed a causal interpretation is investigated. As in Section 3, endogeneity of high school choices is controlled for by explicitly modelling their determinants and by allowing any residual component to freely correlate with the outcomes of interest. Endogenous selection issues are also investigated; however - differently from Section 3 - in this section, a double selectivity process is modelled: selection into job search activity (potentially endogenous for the estimation of employment probabilities) and selection into employment (potentially endogenous for the estimation of low pay probabilities). The model is estimated without taking university attendance into account: see the discussion at the start of Section 3 for the reasons motivating this choice.

### 4.1 The model of early labour market outcomes

The probability of graduating from certain types of high school is modelled in the same fashion of Section 3, but on a partly different sample (see Section 2 for the differences in
sample selection rules in this and the previous sections); thence, the first two equation of the model are equations (1) and (2) of Section 3.

The first labour market outcome considered is the job search decision. Let $r^{*}{ }_{i}$, the latent propensity to engage in job search activities after high school graduation for individual $i$, be a linear function of personal attributes $z_{r}$, high school choices and an error term, assumed to be distributed as standard normal:

$$
\begin{equation*}
r_{i}^{*}=\beta_{r}{ }^{\prime} z_{r i}+\eta_{g} g_{i}+\eta_{p} p_{i}+\varepsilon_{r i} \tag{6}
\end{equation*}
$$

and let $r_{i}=\mathrm{I}\left(r_{i}^{*}>0\right)$ be a dichotomous indicator of the search activity.
The employment status at the date of interview can be observed for those individuals who search for a job (recall from Section 2 that self-employment is excluded from the analysis and that seasonal or occasional jobs are considered as unemployment). Assume that it represents some unobserved employment propensity $e^{*}$, which depends upon personal attributes $m_{e}$ and school choices:

$$
\begin{equation*}
e_{i}^{*}=\beta_{e}^{\prime} m_{e i}+\theta_{g} g_{i}+\theta_{p} p_{i}+\varepsilon_{e i} \quad \text { if } r_{i}=1 \tag{7}
\end{equation*}
$$

and let $e_{i}=\mathrm{I}\left(e_{i}^{*}>0\right)$ indicate whether an individual is employed or not at the interview date. The employment process is censored for those who do not participate in the labour market after graduation.

Besides employment probabilities, another relevant dimension of the school-to-work transition is the wage level that individuals obtain on entry into the labour market. Clearly, entry levels are not informative of future earnings prospects, since the correlation between initial and later earnings varies depending upon which model better describes labour market functioning. The focus is placed on the probability of low pay for two reasons. First, from a substantive point of view, in recent years low pay has emerged as a policy issue in many countries, and research on this topic stresses that as long as there are discontinuities in the earnings process across the distribution, the relevance of earnings determinants might vary according to the earnings quantiles considered. Second, from a modelling point of view, estimating an earnings equation in the multivariate probit context of this paper would have required to assume normality or log-normality of earnings, whereas by focussing on low pay probabilities, normality is required only up to any unspecified transformation of
earnings (as will be shown below). Finally, it has been chosen to derive the low pay cut-off from the sample earnings distribution, i.e. to focus on individual earnings relative to the cohort of high school graduates.

Let individual earnings $w_{i}$ depend upon personal and job attributes $q_{l}$ and high school choices according to the following equation:

$$
\begin{equation*}
f\left(w_{i}\right)=\psi^{\prime} q_{l i}+\omega_{g} g_{i}+\omega_{p} p_{i}+v_{i} \text { if } e_{i}=1 \tag{8.a}
\end{equation*}
$$

where $f()$ is a monotonic unspecified transformation such as the error term $v_{i}$ is distributed as standard normal; let also $p_{25}$ indicate the first quartile of the earnings distribution. Note that the earnings is censored upon employment (and, a fortiori, upon search). An individual is counted as low paid whenever $w_{i}<p_{25}$, i.e. when $f\left(w_{i}\right)<f\left(p_{25}\right)$. By subtracting each side of (8.a) from $f\left(p_{25}\right)$, the earnings process can be rewritten as:

$$
\begin{equation*}
l_{i}^{*}=\beta_{l}^{\prime} q_{l i}+\varphi_{g} g_{i}+\varphi_{p} p_{i}+\varepsilon_{l i} \text { if } e_{i}=1 \tag{8.b}
\end{equation*}
$$

where $l^{*} \equiv f\left(p_{25}\right)-f\left(w_{i}\right)$, the first element of $\beta_{l}$ subsumes the difference between $f\left(p_{25}\right)$ and the constant term in $\psi$, remaining elements of $\beta_{l}$ are the same as the corresponding elements of $\psi$ but with opposite sign, $\varphi_{g, p} \equiv-\omega_{g, p}$ and $\varepsilon_{l i} \equiv-v_{i}$. Note that now coefficients associated to covariates parameterise low wages. Observing an individual in low pay means that the corresponding latent variable $l^{*}{ }_{i}$ is positive; let $l_{i}=\mathrm{I}\left(l^{*}{ }_{i}>0\right)$ be a dichotomous indicator of the low pay event.

As in the case of academic performance, endogeneity of high school choices and (double nested) selection is tackled by means of a multivariate probit model, which emerges if one assumes that the vector of error terms $\varepsilon_{i}^{L}=\left(\varepsilon_{g i}, \varepsilon_{p i}, \varepsilon_{s i}, \varepsilon_{e i}, \varepsilon_{l i}\right)$ follows the five-variate normal distribution:

$$
\begin{equation*}
\varepsilon_{i}^{L} \sim \aleph_{5}(0, \Sigma) \tag{9}
\end{equation*}
$$

where the correlation matrix $\Sigma$ has unit diagonal elements and extra-diagonal elements equal to $\lambda_{j k}, j, k=g, p, r, e, l$. Overall, the model specified is a five-variate probit with nested censoring of two equations. Test of the endogeneity of high school choices and selection
can be performed by testing the significance of the cross-equation correlation coefficients of unobservables.

Identifying restrictions for the model of early labour market outcomes are as follows. The job search equation is identified using the same indicators of family characteristics relevant at the time high school choices were made used in Section 3, i.e. a dummy for having at least one grandparent holding a university degree in the general high school equation, and that dummy plus a dummy indicating if the mother was retired when the interviewed was 14 in the private school equation. Variables that are -plausibly-- not taken into account by employers when screening job candidates might serve for the purpose of identifying the employment equation, the mark reported at junior high school being an example of such variables. Empirically, however, it was found that not only junior high marks, but also high school marks were insignificant in shifting employment probabilities, suggesting that employers do not value the informative content of marks, but probably limit themselves at considering signal embedded in the type of diploma (general versus technical) obtained. Therefore, both marks have been used as instruments for employment probabilities. Finally, indicators of parental backgrounds could be excluded from the low pay equation after their impact on schooling choices and employment probabilities have been controlled for. In practice, it was found that maternal background (both in terms of educational attainment and occupation) produced the better performance in terms of excludability from the low pay equation, so that are these variables the ones used as instruments.

### 4.2 Results

Table 8 illustrates results obtained estimating the economic performance model. For the sake of compactness, high school equations are not reported; results on the determinants of high school choices were very similar to those obtained from the academic performance model (see also the reported correlation coefficient between unobservables of high school equations, virtually identical to the one of Table 6), and are available upon request.

The adopted identification strategy is supported by the data - see the $p$-values for instruments significance in the various equations reported at the bottom of the table. Estimated correlations of unobservables indicate that general high school graduates have a job search propensity which is lower than the one of technical school ones. This could be the case if general high school graduates could count on higher non labour incomes even
after controlling for the wide range of family background indicators available in the data. Alternatively, one might argue that these individuals are of a less labour-market-oriented type compared with graduates from technical schools. A negative relationship also emerges between general high schools and employment probabilities. Again, an explanation based on differential inclinations towards the labour market across high school types could rationalise this fact. On the other hand, the correlation between general high school and low pay propensity is non significant. Taking into account correlations involving the private school equation, the only significant coefficient is the one with low pay, indicating that private school graduates have an intrinsically lower earnings potential relative to private school students. Remaining correlation coefficients indicate that also selectivity processes are a vehicle for estimates bias. For example, the unobservables of search and employment propensities are negatively and significantly correlated, indicating that the intrinsic employability of those who search is lower compared to the population of high school graduates. Also, employment and low pay propensities are negatively correlated: had the unemployed found a job, their low pay chances would have been larger than the ones of the employed, suggesting that unemployment is due to lack of demand rather than to high reservation wages. Tests for the significance of groups of correlation coefficients indicate that none of the high schools processes is ignorable, and that the whole correlation structure is statistically significant, both facts that motivate the use of the simultaneous equations model.

Comparing estimated marginal effects between models with endogenous and exogenous high school choices further motivates the employed modelling strategy. For example, the search probability differential between general and technical schools is $-28 \%$ in the exogenous model of column (2). Endogenisation reduces the estimated effect by a half: the causal impact of general high school reduces search probabilities by $13 \%$, whereas the remaining part of the effect estimated in the exogenous model is due to endogenous sorting. The negative effect that persists after endogenisation indicates that general high schools reduce the probability that a random individual engages in search activities after completing high school compared to technical schools, as could be the case if the type of skills acquired or the overall school environment were more market oriented in the latter case. Taking into account the private/public schools divide, the data indicate that the differential is small and not statistically significant. As for the other determinants of search propensities in column (1), the data indicate that the two sets of characteristics that are
relevant in determining high school choices - family background and previous performances at school - still play a role after the impact of high school choices on search probability has been taken into account. For example, search probabilities are larger when there are no siblings, where parents have high educational attainment, or where parents are located on high ranks of the occupational ladder.

The importance of taking selectivity issues into account is confirmed by looking at results on employability. The probability of (stable) employment estimated at the average of personal characteristics is some $5 \%$ larger in the endogenous model compared with the exogenous one. Recalling that the group at risk of employment (i.e those who actually search) have an intrinsic employment propensity lower than the one for the overall sample (i.e $\rho_{r e}<0$ ), the larger prediction can be seen is a consequence of deriving estimates that refer to the overall sample. The marginal effect of general high schools on employment probabilities is negative and drops by $10 \%$ in absolute value when one moves from the exogenous to the endogenous model. This finding is a result of the negative correlation of unobservables between general high schools and employment equations; net of this negative sorting bias, technical and vocational schools graduates still have an employment probability $20 \%$ larger compared to general schools graduates. It should be stressed that the negative selection effect operates in the opposite direction: since the sample involved in search activity has a lower employment probability compared to the overall sample, the impact of factors that negatively influence employment is overestimated in the sample doing search, and we should observe the marginal effect of general high schools to become more negative after accounting for endogenous selection. The data, however, indicate that negative sorting of less employable individuals into the pool of high school graduates is predominant, and simultaneously accounting for the two endogeneity issues reduces the size of the effect. No significant effect on employment probabilities, on the other hand, can be associated to the choice of a private high school. Remaining marginal effects reveal the existence of a $10 \%$ gender gap and of marked regional differences, youths from the south and islands showing an employment probability which is $41.2 \%$ lower than that of otherwise comparable individuals. The data also indicate the relevance of household characteristics as determinants of employment probabilities, a likely consequence of the association between the former and non-labour incomes on which job searcher can count. For example, the presence of siblings is associated with positive marginal effects, whereas a favourable educational background reduces employment probabilities.

Relevant selection biases emerge if one compares estimates of low pay probabilities from the exogenous and endogenous model, the predicted probability rising from 20 to $34 \%$ moving from the former to the latter: since the sample on which wages are observed (i.e. $e_{i}=1$ ) is a high wage sample (i.e. $\rho_{e}<0$ ), conditioning on employment leads to underestimate low pay probabilities. The same remark applies to the rise in absolute value of the effects of high school choices; in the case of private schools, the rise is also reinforced by making allowance for the positive sorting of low wage individuals into private schools (i.e. $\rho_{p}>0$ ). Overall, the model shows that general high schools significantly increase low pay probabilities relative to technical ones, whereas private school reduce such probabilities compared to public schools. The first finding confirms that the type of education provided by technical and vocational schools increases economic performance. Explanations for the second finding are less evident. Marginal effects for the other covariates show the existence of a relevant gender pay gap and of marked territorial differentials. The father's educational attainment increases low pay probabilities. The mode of search significantly affects low pay chances. In particular, those who have found the current job through a family network (i.e. the excluded category in the estimation of coefficients underlying marginal effects) seem to have the largest low pay probability. Since family networks are the most important search mode in Italy, especially for first job seekers, this result is rather interesting and is consistent with two not mutually exclusive explanation: family networks are used by less able individuals or family networks tend to provide low wage jobs. The other 'wage determinants' included in the low pay equation have marginal effects that go in the expected directions; for example, low pay probabilities decrease as we move towards high level occupations or large firms, and as more tenured employees are taken into account. It should be noted that the absolute value of the effects is larger in the endogenous model compared with the exogenous one, further confirming the importance of allowing for the various selectivity issues analysed in this paper.

## 5. Concluding remarks

The high school choices of young Italians largely depend on two factors: ability and family background. Using indicators of earlier school performance, this paper has shown that holding parental background fixed-better performance individuals tend to select into general high schools. On the other hand, favourable educational and occupational family background push individuals into general high schools, whereas the probability of
graduating from a private school rises with the availability of resources in the family of origin but decreases with school performance prior to high school.

This paper has provided evidence on the effects of high school choices on subsequent academic performances. Graduating from a general high school substantially increases the probability of attending university, whereas no shift in such a probability due to private schools can be found in the data after a proper allowance is made for simultaneity issues. Once at university, the speed in passing exams - a crucial performance indicator for today's Italian students-is found to depend positively on general school attendance and negatively on private school one. Therefore, while attending a general high school can be seen as a way of improving ones academic outcomes, the data indicate that private schools, if anything, reduce them.

High school effects are not confined to college outcomes, but also spread on the school-to-work transition. This time, patterns are reversed, and graduates from technical or vocational schools score much better on either the probability of finding a stable job, or the probability of finding a job above a pre-determined low pay cut-off. Given that the model controls for observed and unobserved heterogeneity, these results are unlikely to reflect a low quality of general schools students that go on the labour market; rather, they may be given a causal interpretation, suggesting that generalist skills are not valued in the youths' labour market. On the other hand, private school attendance is found to reduce low pay probabilities.

From the modelling point of view, this paper has placed considerable emphasis on simultaneity issues, coherently with some previous contributions in this literature. With the aim of tackling those issues, multivariate probability models with partial observability have been developed. Results support the adopted modelling strategy, as both high school choices and selectivity issues matter for the estimation of the effects of interest. Ignoring those issues would have led to wrong conclusions in most cases.

The picture emerging from this paper extends previous findings for Italy. Family background plays a central role in determining school choices, while choices have relevant effects on subsequent educational and economic paths. By deepening the separation between generalist and vocational tracks, the reform of secondary high education which will be implemented in the near future might have the effect of increasing the role of parental backgrounds in shaping individuals' lives. An effective functioning of the mechanisms designed in order to guarantee the 'equal dignity' of tracks -such as the
possibility of track changes after the initial choice- appear as a crucial feature of the implementation phase for preventing intergenerational persistence and social segmentation to increase. As for the other topical dimension of the school debate, i.e. school provision, the negative effect of private schools estimated in the academic performance equation indicates that, on average, private schools are less effective than public ones in enhancing subsequent educational outcomes.

## References

Altonji, G.J., Elder, T.E. and Taber, C.R. (2000): Selection on Observed and Unobserved Variables: Assessing the Effectiveness of Catholic School, National Bureau of Economic Research, Working Paper 7831.
Bertola,G. and Checchi,D. (2001): Sorting and Private Education in Italy, Lavoro e Relazioni Industriali, 2, pp. 87-124.
Brunello, G. and Giannini, M. (2000): Stratified or Comprehensive? The Economic Efficiency of School Design, Institute for Social and Economic Research, University of Essex, Working Paper 32.
Checchi, D., Ichino, A. and Rustichini, A. (1999), More Equal but Less Mobile? Intergenerational Mobility and Inequality in Italy and in the US, Journal of Public Economics, 74 (3), pp. 351-393.
Checchi, D. and Jappelli, T. (2002): School Choice and Quality, unpublished manuscript available at: http://www.economia.unimi.it/checchi/index.html

Evans, W.N. and Schwab, R.M. (1995): Finishing High School and Starting College: Do Catholic Schools Make a Difference?, Quarterly Journal of Economics, 110(4), pp. 941-974.
Neal, D. (1997): The Effects of Catholic Secondary Schooling on Educational Achievement, Journal of Labor Economics, 15(1), pp. 98-123
Figlio, D.N. and Stone, J. (1999): Are Private Schools Really Better? Research in Labor Economics, pp. 115-140.
Figlio, D.N. and Ludwig, J. (2000): Sex, Drugs, and Catholic Schools: Private Schooling and Non-Market Adolescent Behaviors, National Bureau of Economic Research, Working Paper 7990.
Margolis, D.N. and Simonnet, V. (2003): Educational Track, Networks and Labor Market Outcomes, Institute for the Study of Labor, Discussion Paper 699.
Moenjak, T. and Worsick, C. (2003): Vocational Education in Thailand: A Study of Choice and Returns, Economics of Education Review, 22, pp.99-107.

Table 1: The distribution of high school degree types

|  | Private school? | No | Yes | Total | Percent |
| :---: | :---: | :---: | :---: | :---: | :---: |
| General school? |  |  |  |  |  |
| No |  |  |  |  |  |
| Number of cases |  | 9574 | 972 | 10546 | 68.87 |
| Row frequencies |  | 90.78 | 9.22 |  |  |
| Column frequencies |  | 71.30 | 51.51 |  |  |
| Yes |  |  |  |  |  |
| Number of cases |  | 3853 | 915 | 4768 | 31.13 |
| Row frequencies |  | 80.81 | 19.19 |  |  |
| Column frequencies |  | 28.70 | 48.49 |  |  |
| Total |  | 13427 | 1887 | 15314 |  |
| Percent |  | 87.68 | 12.32 |  |  |

Table 2: High school degree types and parental background

|  | Probability of graduating in a <br> general school |  |  | Probability of graduating in a <br> private school |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | By father's <br> education | By mother's <br> education |  | By father's <br> education | By mother's <br> education |
| None or <br> elementary | 14.06 | 14.84 |  | 8.67 | 8.29 |
| Junior high school | 21.97 | 23.02 |  | 11.28 | 11.62 |
| High school | 46.57 | 51.21 |  | 15.04 | 17.15 |
| University degree | 78.97 | 81.76 |  | 18.33 | 15.31 |

Table 3: High school degree types and parental occupation

|  | Probability of graduating in a general school |  | Probability of graduating in a private school |  |
| :---: | :---: | :---: | :---: | :---: |
|  | By father's occupation | By mother's occupation | By father's occupation | By mother's occupation |
| Self employed shop seller/retailer | 32.81 | 30.28 | 22.46 | 21.71 |
| Craft | 21.92 | 24.20 | 11.88 | 16.89 |
| Farmer | 13.40 | 13.41 | 8.17 | 6.15 |
| Entrepreneur | 32.86 | 32.26 | 22.70 | 22.58 |
| Professional | 53.75 | 52.38 | 20.30 | 20.00 |
| Manager | 66.29 | 79.31 | 18.79 | 18.62 |
| Teacher | 68.00 | 67.60 | 6.50 | 13.31 |
| White collar high level | 40.22 | 47.73 | 11.42 | 16.57 |
| White collar low level | 35.18 | 42.04 | 10.77 | 15.45 |
| Blue collar high level | 15.62 | 14.34 | 8.20 | 10.80 |
| Blue collar low level. | 15.45 | 18.17 | 7.87 | 9.87 |
| Not available/ does not know | 27.99 | 24.37 | 14.15 | 10.74 |

Table 4: High school degree types and junior high school marks

|  | Probability of graduating in <br> a general school | Probability of graduating in <br> a private school |
| :--- | :---: | :---: |
| sufficiente (D) | 10.17 | 13.81 |
| buono (C) | 23.63 | 12.10 |
| distinto (B) | 43.56 | 11.32 |
| ottimo (A) | 68.84 | 9.23 |

Table 5: High school degree types and academic performance

|  | Probability of attending <br> university |  | Probability of taking <br> exams quickly* |
| :--- | :---: | :---: | :---: |
| Whole sample | 40.81 | 45.64 |  |
| Graduated in istituti <br> tecnici professionali | 22.34 | 38.62 |  |
| Graduated in a liceo | 81.67 | 49.88 |  |
| Graduated in a public <br> school | 40.02 | 46.13 |  |
| Graduated in a private <br> school | 46.42 | 42.68 |  |

*conditional on university attendance

Table 6: High school degree types and early labour market outcomes

|  | Probability of job search after graduation | Probability of being employed* | Probability of being low paid** |
| :---: | :---: | :---: | :---: |
| Whole sample | 67.71 | 42.77 | 24.76 |
| Graduated in istituti tecnici professionali | 80.95 | 49.18 | 24.23 |
| Graduated in a liceo | 40.47 | 16.40 | 31.31 |
| Graduated in a public school | 68.41 | 43.24 | 25.17 |
| Graduated in a private school | 62.65 | 39.01 | 21.13 |

*Conditional on job search, excludes the self employed
**Conditional on being employed. Low pay defined as bottom quartile of the sample distribution of hourly net pay

Table 7: Models of academic performance

|  | (1) Endogenous choices and selection |  |  |  |  |  |  |  | (2) Exogenous choices and selection |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | P (general) |  | P(private) |  | P (university) |  | P (fast) |  | P (university) |  | P (fast) |  |
| Prediction at sample means | 0. 254 |  | 0.106 |  | 0.394 |  | 0.377 |  | 0.396 |  | 0.451 |  |
| General school |  |  |  |  | 0.429 | (5.97) | 0.153 | (1.77) | 0.469 | (39.71) | 0.084 | (5.22) |
| Private school |  |  |  |  | -0.049 | (0.79) | -0.196 | (2.39) | -0.040 | (2.65) | -0.061 | (3.07) |
| Female | 0.096 | (11.59) | 0.031 | (6.03) | -0.031 | (2.68) | 0.027 | (1.79) | -0.036 | (3.59) | 0.031 | (2.24) |
| Resides in the north east | -0.013 | (1.00) | -0.019 | (2.70) | 0.001 | (0.08) | -0.013 | (0.65) | 0.006 | (0.42) | -0.008 | (0.40) |
| Resides in the centre | 0.004 | (0.34) | -0.056 | (8.55) | 0.015 | (1.00) | -0.152 | (8.18) | 0.015 | (1.06) | -0.156 | (8.26) |
| Resides in the south | 0.010 | (0.81) | -0.039 | (5.68) | 0.064 | (4.40) | -0.242 | (12.73) | 0.065 | (4.53) | -0.263 | (14.18) |
| Has 1 sibling | -0.069 | (5.88) | -0.043 | (6.19) | -0.063 | (4.14) |  |  | -0.061 | (4.20) |  |  |
| Has 2 siblings | -0.081 | (6.20) | -0.051 | (6.81) | -0.083 | (4.77) |  |  | -0.079 | (4.81) |  |  |
| Has 3 siblings | -0.085 | (4.76) | -0.060 | (5.99) | -0.090 | (3.87) |  |  | -0.087 | (3.90) |  |  |
| Has 4 or more siblings | -0.149 | (6.23) | -0.053 | (3.98) | -0.106 | (3.38) |  |  | -0.105 | (3.48) |  |  |
| Father's educational degree |  |  |  |  |  |  |  |  |  |  |  |  |
| Junior high school | 0.034 | (2.65) | 0.016 | (2.06) | 0.003 | (0.21) | 0.001 | (0.05) | 0.003 | (0.18) | -0.005 | (0.20) |
| High school | 0.117 | (7.24) | 0.031 | (3.03) | 0.092 | (4.73) | 0.027 | (0.99) | 0.087 | (4.78) | 0.005 | (0.18) |
| University | 0.295 | (11.26) | 0.061 | (3.81) | 0.225 | (6.40) | 0.051 | (1.34) | 0.212 | (7.12) | 0.013 | (0.39) |
| Not reported | 0.043 | (1.26) | 0.031 | (1.46) | -0.050 | (1.30) | -0.036 | (0.60) | -0.051 | (1.31) | -0.040 | (0.62) |
| Mother's educational degree |  |  |  |  |  |  |  |  |  |  |  |  |
| Junior high school | 0.040 | (3.32) | 0.014 | (1.92) | 0.031 | (2.27) | 0.038 | (1.83) | 0.031 | (2.27) | 0.032 | (1.43) |
| High school | 0.141 | (8.94) | 0.058 | (5.69) | 0.089 | (4.37) | 0.076 | (2.78) | 0.083 | (4.59) | 0.054 | (2.06) |
| University | 0.281 | (9.40) | 0.052 | (2.90) | 0.148 | (3.75) | 0.116 | (2.93) | 0.136 | (3.89) | 0.095 | (2.61) |
| Not reported | 0.163 | (3.86) | 0.051 | (2.03) | -0.028 | (0.58) | -0.030 | (0.43) | -0.032 | (0.69) | -0.037 | (0.50) |
| Father's occupation |  |  |  |  |  |  |  |  |  |  |  |  |
| Self employed shop seller/retailer | 0.127 | (5.77) | 0.130 | (8.31) | 0.105 | (4.03) | 0.093 | (2.44) | 0.099 | (4.07) | 0.060 | (1.61) |
| Craft | 0.045 | (2.29) | 0.035 | (2.69) | 0.056 | (2.52) | 0.065 | (1.89) | 0.054 | (2.44) | 0.055 | (1.49) |
| Farmer | -0.013 | (0.48) | 0.016 | (0.94) | 0.046 | (1.57) | 0.039 | (0.82) | 0.051 | (1.72) | 0.032 | (0.63) |
| Entrepreneur | 0.088 | (3.44) | 0.136 | (7.44) | 0.129 | (4.31) | 0.107 | (2.54) | 0.122 | (4.29) | 0.070 | (1.66) |
| Professional | 0.127 | (5.17) | 0.079 | (4.80) | 0.151 | (5.14) | 0.087 | (2.29) | 0.148 | (5.23) | 0.051 | (1.36) |
| Manager | 0.148 | (6.15) | 0.058 | (3.77) | 0.130 | (4.45) | 0.093 | (2.62) | 0.122 | (4.40) | 0.068 | (1.89) |
| Teacher | 0.065 | (2.07) | -0.045 | (2.53) | 0.103 | (2.58) | 0.052 | (1.24) | 0.096 | (2.42) | 0.049 | (1.12) |
| White collar high level | 0.075 | (4.16) | 0.008 | (0.66) | 0.107 | (5.16) | 0.082 | (2.68) | 0.102 | (4.97) | 0.069 | (2.16) |
| White collar low level | 0.056 | (2.80) | 0.005 | (0.41) | 0.089 | (3.88) | 0.056 | (1.70) | 0.087 | (3.80) | 0.043 | (1.25) |
| Blue collar high level | -0.033 | (2.10) | -0.005 | (0.51) | 0.010 | (0.56) | 0.037 | (1.25) | 0.010 | (0.57) | 0.038 | (1.22) |
| Not reported | 0.025 | (0.75) | 0.022 | (1.08) | 0.108 | (2.85) | 0.078 | (1.43) | 0.103 | (2.72) | 0.058 | (1.01) |




Note: The model in column (1) is estimated via simulated maximum likelihood, using a GHK simulator with 130 random draws. The models in column (2) are univariate probits. The table reports marginal effects evaluated at the mean of explanatory variables. Asymptotic t-ratios in brackets refer to underlying coefficients. P-values from hypotheses tests are reported, degrees of freedom in square brackets. The excluded category is: technical or vocational high school, public high school, male, resides in the north west, has no siblings, has both parents with no or elementary degree in low level manual occupations, reported a mark of D at junior high school, enrolled at high school at 14, is born in 1976, reported a mark between 36 and 40 at high school, has never failed at high school, has never changed high school, has no grandparents with university degree and his/her mother was not retired when the individual was 14 .

Table 8: Models of early labour market outcomes

|  | (1) Endogenous choices and selection |  |  |  |  |  | (2) Exogenous choices and selection |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | P (search) |  | P (employment) |  | P (low pay) |  | P (search) |  | P (employment) |  | P (low pay) |  |
| Prediction at sample means | 0.707 |  | 0.454 |  | 0.338 |  | 0.708 |  | 0.405 |  | 0.198 |  |
| General school | -0.130 | (2.03) | -0.197 | (4.20) | 0.176 | (2.02) | -0.280 | (26.14) | -0.293 | (20.3) | 0.060 | (2.24) |
| Private school | 0.045 | (0.76) | -0.011 | (0.15) | -0.166 | (2.07) | 0.019 | (1.46) | 0.003 | (0.14) | -0.030 | (1.35) |
| Female | 0.049 | (4.44) | -0.096 | (8.12) | 0.156 | (8.31) | 0.062 | (7.22) | -0.087 | (7.94) | 0.107 | (7.17) |
| Resides in the north east | 0.004 | (0.31) | 0.057 | (3.35) | -0.079 | (3.55) | 0.022 | (0.17) | 0.056 | (3.30) | -0.050 | (2.78) |
| Resides in the centre | 0.017 | (1.35) | -0.174 | (11.10) | 0.138 | (4.88) | 0.015 | (1.25) | -0.169 | (11.35) | 0.080 | (4.22) |
| Resides in the south | 0.015 | (1.23) | -0.412 | (26.75) | 0.414 | (9.44) | 0.013 | (1.09) | -0.400 | (27.79) | 0.265 | (11.15) |
| Has 1 sibling | 0.063 | (4.88) | 0.031 | (1.77) | -0.011 | (0.45) | 0.054 | (4.44) | 0.036 | (2.11) | 0.005 | (0.22) |
| Has 2 siblings | 0.088 | (5.97) | 0.037 | (1.79) | 0.026 | (0.88) | 0.075 | (5.53) | 0.045 | (2.29) | 0.042 | (1.76) |
| Has 3 siblings | 0.080 | (4.09) | 0.038 | (1.41) | 0.025 | (0.61) | 0.067 | (3.56) | 0.043 | (1.65) | 0.041 | (1.21) |
| Has 4 or more siblings | 0.114 | (4.28) | 0.021 | (0.62) | 0.050 | (0.98) | 0.095 | (3.70) | 0.028 | (0.86) | 0.061 | (1.41) |
| Father's educational degree |  |  |  |  |  |  |  |  |  |  |  |  |
| Junior high school | 0.0001 | (0.01) | -0.011 | (0.71) | 0.025 | (1.28) | 0.005 | (0.38) | -0.007 | (0.50) | 0.014 | (0.88) |
| High school | -0.049 | (2.78) | -0.041 | (1.96) | 0.058 | (1.94) | -0.033 | (2.05) | -0.039 | (1.93) | 0.027 | (1.16) |
| University | -0.169 | (5.70) | -0.120 | (2.95) | 0.159 | (2.06) | -0.130 | (5.34) | -0.122 | (3.20) | 0.072 | (1.17) |
| Not reported | -0.021 | (0.62) | 0.005 | (0.10) | 0.167 | (2.95) | -0.018 | (0.52) | 0.011 | (0.25) | 0.145 | (2.95) |
| Mother's educational degree |  |  |  |  |  |  |  |  |  |  |  |  |
| Junior high school | -0.005 | (0.40) | -0.023 | (1.65) |  |  | 0.0001 | (0.04) | -0.022 | (1.59) |  |  |
| High school | -0.070 | (3.85) | -0.035 | (1.63) |  |  | -0.048 | (3.09) | -0.033 | (1.62) |  |  |
| University | -0.159 | (5.06) | -0.056 | (1.13) |  |  | -0.121 | (4.41) | -0.073 | (1.52) |  |  |
| Not reported | -0.108 | (2.49) | 0.005 | (0.09) |  |  | -0.088 | (2.07) | -0.018 | (0.34) |  |  |
| Father's occupation |  |  |  |  |  |  |  |  |  |  |  |  |
| Self employed shop seller/retailer | -0.145 | (5.93) | -0.002 | (0.07) | -0.009 | (0.21) | -0.128 | (5.65) | -0.014 | (0.51) | -0.024 | (0.77) |
| Craft | -0.098 | (4.75) | 0.015 | (0.62) | 0.005 | (0.16) | -0.093 | (4.49) | 0.006 | (0.27) | 0.004 | (0.16) |
| Farmer | -0.064 | (2.35) | -0.014 | (0.45) | -0.079 | (1.76) | -0.064 | (2.32) | -0.022 | (0.74) | -0.069 | (2.05) |
| Entrepreneur | -0.161 | (5.73) | 0.073 | (2.14) | -0.041 | (0.88) | -0.148 | (5.56) | 0.058 | (1.79) | -0.034 | (0.96) |
| Professional | -0.187 | (6.98) | 0.018 | (0.50) | -0.028 | (0.52) | -0.166 | (6.50) | 0.002 | (0.05) | -0.029 | (0.70) |
| Manager | -0.098 | (3.94) | -0.010 | (0.29) | 0.021 | (0.42) | -0.077 | (3.30) | -0.011 | (0.34) | 0.006 | (0.13) |
| Teacher | -0.083 | (2.65) | -0.029 | (0.54) | -0.001 | (0.01) | -0.077 | (2.48) | -0.033 | (0.62) | -0.005 | (0.06) |
| White collar high level | -0.091 | (4.84) | -0.007 | (0.30) | 0.023 | (0.68) | -0.082 | (4.44) | -0.013 | (0.59) | 0.012 | (0.45) |
| White collar low level | -0.099 | (4.83) | 0.012 | (0.48) | -0.023 | (0.63) | -0.093 | (4.51) | 0.002 | (0.08) | -0.020 | (0.69) |
| Blue collar high level | -0.021 | (1.30) | 0.060 | (3.37) | -0.046 | (1.80) | -0.026 | (1.58) | 0.055 | (3.13) | -0.024 | (1.14) |
| Not reported | -0.140 | (3.97) | -0.046 | (1.07) | -0.127 | (1.85) | -0.136 | (3.86) | -0.060 | (1.44) | -0.111 | (2.31) |


|  | (1) Endogenous choices and selection |  |  |  |  | (2) Exogenous choices and selection |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | P (search) |  | P (employment) |  | P (low pay) | P (search) |  | P (employment) |  | P (low pay) |
| Mother's occupation |  |  |  |  |  |  |  |  |  |  |
| Self employed shop seller/retailer | -0.037 | (1.38) | 0.032 | (0.94) |  | -0.031 | (1.15) | 0.022 | (0.65) |  |
| Craft | 0.000 | (0.01) | 0.086 | (1.82) |  | 0.007 | (0.17) | 0.069 | (1.44) |  |
| Farmer | 0.072 | (1.65) | -0.066 | (1.32) |  | 0.066 | (1.50) | -0.055 | (1.12) |  |
| Entrepreneur | -0.172 | (2.45) | -0.063 | (0.63) |  | -0.168 | (2.39) | -0.085 | (0.86) |  |
| Professional | -0.005 | (0.11) | -0.101 | (1.34) |  | 0.001 | (0.02) | -0.087 | (1.15) |  |
| Manager | -0.083 | (1.71) | -0.148 | (1.68) |  | -0.059 | (1.26) | -0.144 | (1.65) |  |
| Teacher | -0.001 | (0.04) | -0.084 | (2.42) |  | 0.006 | (0.25) | -0.086 | (2.50) |  |
| White collar high level | -0.025 | (1.07) | -0.047 | (1.51) |  | -0.015 | (0.63) | -0.057 | (1.86) |  |
| White collar low level | -0.011 | (0.49) | -0.021 | (0.72) |  | -0.006 | (0.25) | -0.018 | (0.61) |  |
| Blue collar high level | 0.084 | (3.20) | 0.037 | (1.20) |  | 0.079 | (2.98) | 0.050 | (1.62) |  |
| Not reported | -0.002 | (0.11) | 0.004 | (0.23) |  | 0.0001 | (0.01) | 0.002 | (0.09) |  |
| Junior high school mark (D to |  |  |  |  |  |  |  |  |  |  |
| Buono (C) | -0.032 | (2.41) |  |  |  | -0.023 | (1.90) |  |  |  |
| Distinto (B) | -0.084 | (4.05) |  |  |  | -0.051 | (3.69) |  |  |  |
| Ottimo (A) | -0.143 | (4.59) |  |  |  | -0.083 | (5.35) |  |  |  |
| Missing | -0.041 | (1.90) |  |  |  | -0.027 | (1.34) |  |  |  |
| High school enrollment at |  |  |  |  |  |  |  |  |  |  |
| 13 | 0.009 | (0.52) |  |  |  | 0.009 | (0.54) |  |  |  |
| 15 | 0.046 | (2.51) |  |  |  | 0.042 | (2.28) |  |  |  |
| 16 or more | 0.029 | (0.82) |  |  |  | 0.030 | (0.82) |  |  |  |
| Born before 1970 | 0.231 | (2.97) |  |  |  | 0.229 | (2.92) |  |  |  |
| Born between 1970 and 1973 | 0.152 | (4.05) |  |  |  | 0.145 | (3.90) |  |  |  |
| Born in 1974 | 0.118 | (4.58) |  |  |  | 0.113 | (4.31) |  |  |  |
| Born in 1975 | 0.047 | (2.65) |  |  |  | 0.044 | (2.48) |  |  |  |
| Born in 1977 | -0.042 | (2.05) |  |  |  | -0.022 | (1.18) |  |  |  |
| High school mark (36 to 60) |  |  |  |  |  |  |  |  |  |  |
| from 41 to 45 | -0.042 | (3.62) |  |  |  | -0.044 | (3.68) |  |  |  |
| from 46 to 50 | -0.112 | (8.83) |  |  |  | -0.113 | (8.77) |  |  |  |
| from 51 to 55 | -0.163 | (10.2) |  |  |  | -0.166 | (10.28) |  |  |  |
| from 56 to 59 | -0.237 | (11.2) |  |  |  | -0.240 | (11.29) |  |  |  |
| 60 | -0.296 | (13.7) |  |  |  | -0.297 | (13.77) |  |  |  |


| Has failed during high school | (1) Endogenous choices and selection |  |  |  |  | (2) Exogenous choices and selection |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | P (search) |  | P (employment) | P (low pay) |  | P(search) |  | P (employment) | P (low pay) |  |
|  | 0.020 | (1.26) |  |  |  | 0.018 | (1.11) |  |  |  |
| Has changed high school | 0.018 | (0.97) |  |  |  | 0.017 | (0.90) |  |  |  |
| Age at end of high school | -0.020 | (2.61) |  |  |  | -0.020 | (2.60) |  |  |  |
| Search mode current job |  |  |  |  |  |  |  |  |  |  |
| Knew the employer personally |  |  |  | -0.005 | (0.19) |  |  |  | -0.004 | (0.19) |
| Contacted by the employer |  |  |  | -0.104 | (3.57) |  |  |  | -0.083 | (3.70) |
| Job advertisement |  |  |  | -0.053 | (1.93) |  |  |  | -0.042 | (1.89) |
| Sending CVs to employers |  |  |  | -0.052 | (2.44) |  |  |  | -0.043 | (2.46) |
| Public competition |  |  |  | -0.253 | (4.53) |  |  |  | -0.177 | (4.99) |
| Family firm |  |  |  | 0.051 | (1.22) |  |  |  | 0.043 | (1.19) |
| Employment agencies |  |  |  | -0.081 | (2.24) |  |  |  | -0.064 | (2.23) |
| Part-time |  |  |  | -0.171 | (6.40) |  |  |  | -0.131 | (7.56) |
| Work and training contract |  |  |  | 0.058 | (2.31) |  |  |  | 0.050 | (2.43) |
| Fixed term contract |  |  |  | -0.069 | (2.77) |  |  |  | -0.059 | (2.85) |
| Occupation |  |  |  |  |  |  |  |  |  |  |
| White collar high level |  |  |  | -0.134 | (1.66) |  |  |  | -0.108 | (1.80) |
| White collar low level |  |  |  | -0.005 | (0.25) |  |  |  | -0.004 | (0.26) |
| Blue collar high level |  |  |  | -0.003 | (0.14) |  |  |  | -0.003 | (0.18) |
| Employee of family firm |  |  |  | 0.088 | (1.69) |  |  |  | 0.075 | (1.64) |
| Apprenticeship |  |  |  | 0.230 | (5.26) |  |  |  | 0.219 | (5.66) |
| Public sector |  |  |  | 0.009 | (0.35) |  |  |  | 0.008 | (0.34) |
| Industry |  |  |  |  |  |  |  |  |  |  |
| Agriculture |  |  |  | 0.022 | (0.46) |  |  |  | 0.022 | (0.53) |
| Retail trade |  |  |  | 0.035 | (1.61) |  |  |  | 0.026 | (1.41) |
| Transport \& communication |  |  |  | 0.004 | (0.11) |  |  |  | 0.000 | (0.01) |
| Financial services |  |  |  | -0.131 | (2.51) |  |  |  | -0.103 | (2.64) |
| Public administration, |  |  |  | -0.018 | (0.47) |  |  |  | -0.019 | (0.61) |
| Education, Health |  |  |  |  |  |  |  |  |  |  |
| Housing or IT services |  |  |  | 0.079 | (2.07) |  |  |  | 0.067 | (2.01) |
| Other services |  |  |  | 0.105 | (3.88) |  |  |  | 0.092 | (3.95) |



|  | (1) Endogenous choices and selection | (2) Exogenous choices and selection |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Number of observations | 14420 | 14420 | 9764 |  |  |
| Log likelihood | -25686.232 | -7261.268 | -5513.3071 | -1868 |  |
| Model's chi2 | 0.0000 [258] | 0.0000 [60] | 0.0000 [40] | 0.0000 | [55] |
| Note: The model in column The table reports marginal are reported, degrees of fre siblings, has both parents w reported a mark between 36 was not retired when the in firm, started current job in | lated maximum likelihood, using a GHK si mean of explanatory variables. Asymptotic t The excluded category is: technical or voc ree in low level manual occupations, reporte as never failed at high school, has never ch current job via family networks, works full | random draws. Th refer to underlyin ool, public high unior high school, , has no grandpar ent contract, wo | dels in column (2) fficients. P-values male, resides in led at high school ith university degr private sector man | variate prob ypotheses h west, h born in his/her m ing for a |  |

## Quaderni dell'Istituto di Economia dell'Impresa e del Lavoro

1. Solimene L., Market Failures and State Intervention
2. Solimene L., The Efficiency under Private and Public Ownership: Some Empirical Evidence
3. Baici E., Dell'Aringa C., The EMS Effect on the Italian Labour Market
4. Lucifora C., Union Density and Relative Wages: Is there a Relationship?
5. Lucifora C., Sestito P., Determinazione del salario in Italia: una rassegna della letteratura empirica
6. Martini G., Testing Different Bargaining Theories: A Pilot Experiment
7. Lucifora C., Rappelli F., Profili retributivi e carriere:un'analisi su dati longitudinali
8. Dell'Aringa C., Lucifora C., Wage Dispersion and Unionism: Are Unions Egalitarian?
9. Martini G., Horizontal Price Fixing and Antitrust Policy: A Sequentially Rational Design
10. Cassuti G., Dell'Aringa C., Lucifora C., Labour Turnover and Unionism
11. Solimene L., Regolamentazione ed incentivi all'innovazione nel settore delle telecomunicazioni
12. Bigard A., Guillotin Y., Lucifora C. e F. Rappelli, An International Comparison of Earnings Mobility: The Case of Italy and France
13. Martini G., Laboratory Tests of a Kinked Demand Curve Model with Discounting and Game-theoretic Foundations
14. Martini G., A Multi-period Antitrust Game: The Dynamic Effects of Competition Policy
15. Piccirilli G., Monetary Business Cycles with Imperfect Competition and Endogenous Growth
16. Dell'Aringa C., Pay Determination in the Public Service: An International Comparison
17. Lucifora C., Rules Versus Bargaining: Pay Determination in the Italian Public Sector
18. Piccirilli G., Hours and Employment in a Stochastic Model of the Firm
19. Cappellari L., The Covariance Structure of Italian Male Wages, 1974-1988
20. Lucifora C., Working Poors? An Analysis of Low Wage Employment in Italy
21. Lucifora C., Origo F., Alla ricerca della flessibilità: un'analisi della curva dei salari in Italia
22. Dell'Aringa C., Vignocchi C., Employment and Wage Determination for Municipal Workers: The Italian Case
23. Cappellari L., Wage Inequality Dynamics in the Italian Labour Market: Permanent Changes or Transitory Fluctuations?
24. Cappellari L., Low-pay transitions and attrition bias in Italy: a simulated maximum likelihood approach
25. Pontarollo E., Vitali F., La gestione del parco tecnologico elettromedicale tra outsourcing e integrazione verticale
26. Cappellari L., Do the 'Working Poors' Stay Poor? An Analysis of Low-Pay Dynamics in Italy
27. Dell'Aringa C., Lucifora C., Inside the black box: labour market institutions, wage formation and unemployment in Italy
28. Filippini L., Martini G., Vertical Differentiation and Innovation Adoption
29. Lucifora C., Simmons R., Superstar Effects in Italian Football: an Empirical Analysis
30. Brunello G., Lucifora C., Winter-Ebmer R., The Wage Expectations of European College Students
31. Cappellari L., Earnings dynamic and uncertainty in Italy: How do they differ between the private and public sectors?
32. Piccirilli G., Unions and Workforce Adjustment Costs
33. Dell'Aringa C., The Italian Labour Market: Problems and Prospects
34. Bryson A., Cappellari L., Lucifora C., Does Union Membership Really Reduce Job Satisfaction?
35. Cappellari L., The effects of high school choices on academic performance and early labour market outcomes

I paper sono disponibili presso:
Papers are available at:

Istituto di Economia dell'Impresa e del Lavoro
Università Cattolica del Sacro Cuore
Largo Gemelli, 1
20123 Milano (ITALY)
e-mail: ist.eil@unicatt.it
www.unicatt.it/Ist_Ec_Impresa_e_Lavoro


[^0]:    ${ }^{*}$ Helpful conversations with Carlo Dell'Aringa and Claudio Lucifora and data availability from ISTAT through ISFOL are gratefully acknowledged.

[^1]:    ${ }^{1}$ The reform goes in the opposite direction compared to the one introduced by the previous government in the late 1990s, in which track choice was deferred to the third year of secondary high education, after two years

[^2]:    of uniform education.

[^3]:    ${ }^{2}$ Before 1968 graduates from technical or professional high schools were required to pass an additional exam before they could enrol into university.

[^4]:    ${ }^{3}$ Computation of multi-variate normal distributions is performed by simulation, in particular by applying the so-called GHK simulator.

[^5]:    ${ }^{4}$ Parental labour market status when interviewees were 14 is available in the data, and tests indicate that, as a whole, it can be excluded from high school equations. Taken in isolation, the dummy for mother retired is

